# How to Use the Internet for Election Surveys 

## Simon Jackman and Douglas Rivers

Stanford University and Polimetrix, Inc.
May 9, 2008

## Theory and Practice

|  |  | Theory |  |
| :---: | ---: | :---: | :---: |
| Practice | Works | Works | Great! |
| Doesn't work |  |  |  |
| Black magic |  |  |  |

## What Works in Theory and Practice

■ Current Population Survey
■ Biennial Registration and Voting Supplements in November of election years
■ High response rate ( $92 \%$ in 2004)
■ Overstates registration and voting slightly (about 2-3\%)

- Aside from misreporting, it can't have large errors

■ By method of Cochran, Mosteller and Tukey, error bound is about 7\%

## 2004 National Election Study

■ $66 \%$ response rate in pre-election wave with $88 \%$ reinterview rate for an overall response rate of $58 \%$

- $89 \%$ registration rate and $76 \%$ turnout rate are too high.

■ Versus CPS estimates of $72 \%$ registration and $64 \%$ turnout

- Actual was about 70\% registration and 61\% VEP turnout

■ CMT bounds for NES are typically around $\pm 30 \%$
■ But, in practice, it works reasonably well.

## RDD Telephone Surveys

■ Actual response rates for short field periods are very low (15-20\%)
■ Higher contact (and response rates) possible for longer field periods
■ Within-household selection usually non-random

- Gallup uses quotas
- ABC/Washington Post uses unequal probabilities of selection, but ignores in weighting
■ Oldest male/youngest female-something that fails in theory, but seems to work in practice


## Errors in Democratic Primary Polls in 2008

Poll Errors


## Actual and Theoretical Error Distribution in NH

Distribution of Poll Errors


## Don't believe the MOE!

Sample Size and Reported Margin of Error


## Incorporating a Design Effect for Weighting

Assuming Design Effect of 2.56


## 2008 Web Election Surveys

- 2008 is the first year with large Internet-based election surveys

■ NES (dedicated panel, recruited by KN)

- NAES and AP/Yahoo/Harvard (KN access panel)

■ CCAP, CCES, and campaigns (YG/Polimetrix access panel)

## Benefits of RDD Recruitment

■ Coverage of non-Internet households
■ RDD widely accepted as "probability-based"

- Can calculate a meaningful response rate

■ Avoids "volunteers"

## Problems with RDD Recruitment

- Expensive
- Small
- Overuse

■ Response rates are low
■ Attrition is high
■ Practical fixes require abandoning theoretical purity

## An Alternative Approach

- Sample selection from a panel with unknown selection probabilities

■ Use large opt-in panel

- Availability of large amounts of auxiliary data on both population and panel from voter and consumer databases, high quality probability samples

■ We know a lot about both the target population and the panelists and should use it.
■ Sample matching used for selection of subsamples from the panel

■ Works in both theory and practice

## Comparison of RDD with Web Opt-ins

| Group | Unweighted <br> RDD | Web <br> Opt-ins | 2004 <br> ACS | 2003 CPS <br> Internet |
| :--- | :---: | :---: | :---: | :---: |
| Blacks | $7.9 \%$ | $4.3 \%$ | $11.8 \%$ | $9.3 \%$ |
| Hispanics | $4.8 \%$ | $3.3 \%$ | $13.5 \%$ | $7.2 \%$ |
| Postgrad | $17.2 \%$ | $23.3 \%$ | $9.4 \%$ | $14.7 \%$ |
| Age 18-24 | $6.4 \%$ | $8.7 \%$ | $10.3 \%$ | $16.0 \%$ |
| Male | $41.9 \%$ | $58.8 \%$ | $48.9 \%$ | $48.7 \%$ |
| Married | $57.7 \%$ | $60.4 \%$ | $54.3 \%$ | $55.3 \%$ |

## Bias and SEs

■ Standard errors measure sampling variability, not bias.
■ Possible to calculate SEs without knowing data generating process

- Observations are, at a minimum, exchangeable and usually independent.
■ No logical difference between nonresponse and self-selection.

■ In both cases, the selection probabilities are unknown

- Validity of estimates depends upon untestable modeling assumptions
■ Standard approach for both RDD and Web panels leave substantial amounts of bias


## Auxilliary Information

- Voter and consumer databases provide a sampling frame for social science research.
- Frame contains a large amount of auxiliary information that can be used for bias reduction.
- Age, gender, vote history, address, name
- Home value, children, interests, magazines

■ Data available for everyone-both panelists and population.

Idea: Draw a sample from the frame and find the closest matching respondents from a panel.

## Sample Matching

■ Recruit a large reservoir of respondents who are accessible for interviewing (the "panel").
■ Obtain a population frame containing auxiliary information for matching.
■ Select a target sample from the frame.
■ For each unit in the target sample, find the closest matching unit in the reservoir. This is the matched sample.

Variants:
■ Use a high quality sample from another source as the target sample.
■ Dynamic matching: match to multiple studies simultaneously using a flow of invitations.

## Notation

$$
\begin{aligned}
N & =\text { size of panel } \\
n & =\text { size of sample }
\end{aligned}
$$

$X=$ covariates ( $k$ vector)
$Y=$ survey measurements
$Z=$ panel membership indicator
Panel

$$
\left(\tilde{X}_{1}, \tilde{Y}_{1}\right), \ldots,\left(\tilde{X}_{N}, \tilde{Y}_{N}\right)
$$

## Distributions and Parameters

$$
\begin{aligned}
& f_{X}(x)=\text { density of } X \text { in population } \\
& \tilde{f}_{X}(x)=\text { density of } X \text { conditional on } Z_{i}=1
\end{aligned}
$$

$$
f_{Y \mid X}(y \mid x)=\text { conditional distribution of } Y \text { given } X
$$

$\tilde{f}_{Y \mid X}(y \mid x)=$ conditional distribution in the panel

$$
\begin{aligned}
\mu(x) & =E(Y \mid X=x)=\int y f_{Y \mid X}(y \mid x) d x \\
\theta_{0} & =E(Y)=\int \mu(x) f_{X}(x) d x \\
\sigma_{1}^{2}(x) & =V(Y \mid X=x, Z=1)
\end{aligned}
$$

## Assumptions

■ IID Data Generating Process $\left(X_{i}, Y_{i}, Z_{i}\right)$ are i.i.d.
■ Ignorable Selection $f_{Y \mid X}(y \mid x)=\tilde{f}_{Y \mid X}(y \mid x)$
■ Continuous Covariates $X$ has a continuous distribution with bounded convex support
■ Overlap The support of $X$ is the same in the panel as in the population with density bounded away from zero
■ Continuity $\mu(x)$ is Lipschitz continuous
■ Regularity $V(Y \mid X, Z=1)$ is uniformly bounded

## Matching Process

■ Target Sample: Choose a (stratified) random sample of size $n$ from the frame $\left(X_{1}, \ldots, X_{n}\right)$.

- For each element of the target sample, find the closest matching element $M(i)$ in the panel:

$$
\begin{aligned}
& \qquad M(i)=j \text { iff }\left|X_{i}-\tilde{X}_{j}\right| \leq\left|X_{i}-\tilde{X}_{\ell}\right| \text { for all } \ell \text { in the panel } \\
& \text { Let } X_{i}^{*}=\tilde{M}(i) \text { and } Y_{i}^{*}=\tilde{Y}_{M(i)}
\end{aligned}
$$

## Matched Sample

$$
\left(X_{1}^{*}, Y_{1}^{*}\right), \ldots,\left(X_{n}^{*}, Y_{n}^{*}\right)
$$

Matching Estimator

$$
\tilde{\theta}=n^{-1} \sum_{i=1}^{n} \tilde{Y}_{i}
$$

## Scalar Matching

$\square$ The conditional density of $X_{i}^{*}$ given $X_{i}=x$ is

$$
N \tilde{f}(x)\left[1-\tilde{F}_{X}\left(x+\left|x^{*}-x\right|\right)+\tilde{F}_{X}\left(x-\left|x^{*}-x\right|\right)\right]^{N-1}
$$

where $\tilde{F}_{X}$ is the distribution function of $X$ in the panel.
■ Conditional on $X_{i}=x$, the limiting distribution of $N\left(X_{i}^{*}-x\right)$ is Laplace with mean zero and variance $1 / 2 \tilde{f}_{X}(x)^{2}$.
■ The approximate distribution of $X_{i}^{*}-X_{i}$ is, thus, Laplace with mean zero and variance $O\left(1 / N^{2}\right)$

## Asymptotic Distribution of $X^{*}-X$



## Theoretical Results

■ When matching on a $\sqrt{n}$-consistent estimator of the propensity score, the matched sample is consistent and asymptotically normal.
■ When using nearest-neighbors matching and the number of matching variables is greater than two, the estimate is consistent, but involves a bias of order $O\left(n^{k / 2} / N\right)$ where $k$ is the number of matching variables.
■ In general, propensity score matching is to be preferred.

## Monte Carlo Simulations

■ Covariates have different means and covariance structure
■ Support of $X$ is $[-1,1] \times[1,1]$

- Population distribution is truncated bivariate normal with mean zero, SD 1, and correlation 0.3
- Panel distribution is truncated bivariate normal with mean (0.2, -0.3), SD 1, and correlation -0.5

■ $Y \mid X \sim N\left(X_{1}+X_{2} / 2,1\right)$ (ignorable selection)
■ Sample size of $n=1000$
■ Panel size of $N=1500,2000,5000$ or 10000.
■ 1000 Monte Carlo Simulations

## Simulated Bias of Estimators



## Simulated RMSE of Estimators



## Empirical Application: $\mathbf{2 0 0 6}$ CCES

■ Consortium of 37 universities

- Pre- and post election interviews

■ $n=37,000$
■ $N=129,000$
■ Frame was 2005 ACS matched to 2004 NEP Exit Poll
■ 7-point party ID and 5-point ideology imputed from 2004 NAES

■ Matched on age, race, gender, education, income, martial status, party ID, and ideology.

## Senate Estimates and 95\% Confidence Intervals



## Governor Estimates and 95\% Confidence Intervals



## Comparison of Accuracy

| Source | $n$ | Mean Error | RMSE |
| :--- | :---: | :---: | :---: |
| Phone | 255 | 2.76 | 8.34 |
| Rasmussen (IVR) | 83 | 3.82 | 8.47 |
| SurveyUSA (IVR) | 63 | 3.4 | 7.25 |
| Zogby (Internet) | 72 | 4.86 | 9.36 |
| Polimetrix (Internet) | 40 | -0.47 | 5.21 |

Source: Blumenthal and Franklin (2007)

## Some Lessons from 2006

■ Some categories are in short supply, even in large panels:
■ Older, low education minorities
■ Panelists recruited through public opinion surveys have high levels of political interest (and over-report registration and turnout)

■ The much-maligned "paid survey takers" are helpful for matching unregistered voters.
■ Matching is imperfect, so weighting is important to remove remaining biases.

- However, much smaller weights are needed after matching.

■ Panel attrition requires modeling regardless of the method of sampling.

■ The question is not whether you do weighting, but whether you do it well.

## Party ID Trends 2007-08

Democratic


Republican


## Political and Campaign Interest 2007-08

Interest in Politics and Public Affairs


Interest in Campaign


## Trial Heats 2008

Obama vs. McCain


Clinton vs. McCain


